

FSAN/ELEG815: Statistical Learning

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7: Lasso Regression

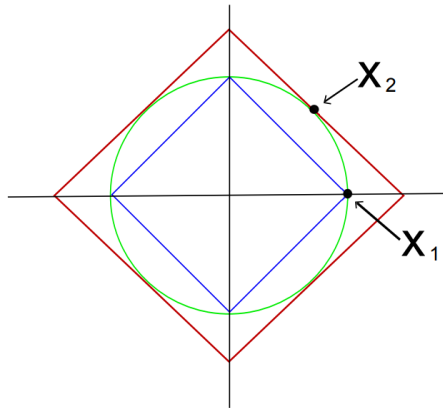
The l_2 Norm and Sparsity

- ▶ The l_0 norm is defined by: $\|\mathbf{x}\|_0 = \#\{i : x(i) \neq 0\}$
Sparsity of $\mathbf{x}$ is measured by its number of non-zero elements
- ▶ The l_1 norm is defined by: $\|\mathbf{x}\|_1 = \sum_i |x(i)|$
 l_1 norm as two key properties:
 - ▶ Robust data fitting
 - ▶ Sparsity inducing norm
- ▶ The l_2 norm is defined by: $\|\mathbf{x}\|_2 = (\sum_i |x(i)|^2)^{1/2}$
 l_2 norm is not effective in measuring sparsity of $\mathbf{x}$

Why l_1 Norm Promotes Sparsity?

Given two N -dimensional signals:

- $x_1 = (1, 0, \dots, 0) \rightarrow$ "Spike" signal
 - $x_2 = (1/\sqrt{N}, 1/\sqrt{N}, \dots, 1/\sqrt{N}) \rightarrow$ "Comb" signal
-
- x_1 and x_2 have the same ℓ_2 norm:
 $\|x_1\|_2 = 1$ and $\|x_2\|_2 = 1$.
 - However, $\|x_1\|_1 = 1$ and $\|x_2\|_1 = \sqrt{N}$.

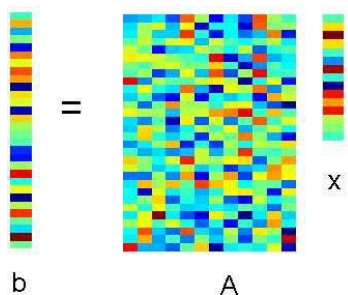


l_1 Norm in Regression

- Linear regression is widely used in science and engineering.

Given $A \in R^{m \times n}$ and $b \in R^m$; $m > n$

Find x s.t. $b = Ax$ (overdetermined)



The diagram illustrates the equation $b = Ax$ using heatmaps. On the left is a vertical column of colored squares labeled b . In the middle is an equals sign. To the right of the equals sign is a square matrix of colored squares labeled A , followed by a vertical column of colored squares labeled x .

ℓ_1 Norm Regression

Two approaches:

- Minimize the ℓ_2 norm of the residuals

$$\min_{x \in R^n} \| \mathbf{b} - \mathbf{A}x \|_2$$

The ℓ_2 norm penalizes large residuals

- Minimizes the ℓ_1 norm of the residuals

$$\min_{x \in R^n} \| \mathbf{b} - \mathbf{A}x \|_1$$

The ℓ_1 norm puts much more weight on small residuals

Matlab Code

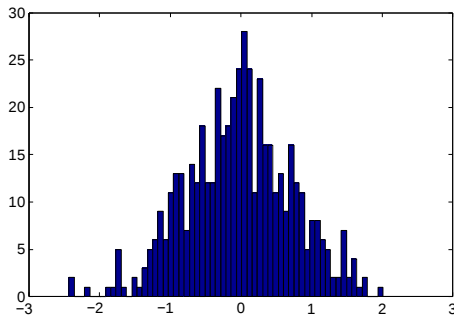
```
min $_{\mathbf{x} \in \mathbb{R}^n}$   $\|\mathbf{b} - \mathbf{Ax}\|_2$   
A=randn(500,150);  
b=randn(500,1);  
x=(A*A)-1 * A * B;
```

Least Squares Solution

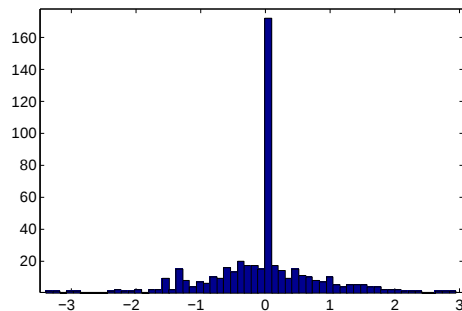
```
min $_{\mathbf{x} \in \mathbb{R}^n}$   $\|\mathbf{b} - \mathbf{Ax}\|_1$   
A=randn(500,150);  
b=randn(500,1);  
X=medrec(b,A,max(A'*b),0,100,1e-5);
```

ℓ_1 Norm Regression

$m = 500, n = 150$. $A = \text{randn}(m, n)$ and $b = \text{randn}(m, 1)$



ℓ_2 Residuals

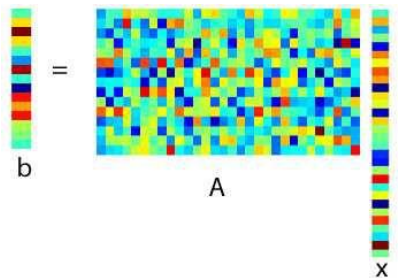


ℓ_1 Residuals

l_1 Norm Regression

Given $A \in R^{m \times n}$ and $b \in R^m$; $m < n$

Find x s.t. $b = Ax$ (underdetermined)



l_1 Norm Regression

Two approaches:

- Minimize the ℓ_2 norm of $\mathbf{x}$

$$\min_{\mathbf{x} \in \mathbb{R}^n} \|\mathbf{x}\|_2 \quad \text{subject to} \quad \mathbf{Ax} = \mathbf{b}$$

- Minimize the ℓ_1 norm of $\mathbf{x}$

$$\min_{\mathbf{x} \in \mathbb{R}^n} \|\mathbf{x}\|_1 \quad \text{subject to} \quad \mathbf{Ax} = \mathbf{b}$$

Matlab Code

- $\min_{\mathbf{x} \in \mathbb{R}^n} \|\mathbf{x}\|_2$ subject to $\mathbf{Ax} = \mathbf{b}$

```
A=randn(150,500);
```

```
b=randn(150,1);
```

```
C=eye(150,500);
```

```
d=zeros(150,1);
```

```
X=lsqlin(C,d,[],[],A,b);
```

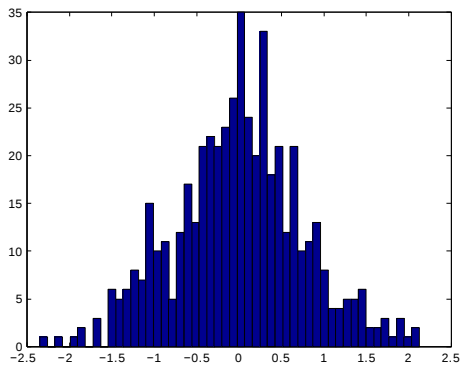
- In general:

$$\min_{\mathbf{x} \in \mathbb{R}^n} f(\mathbf{x}) \quad \text{subject to} \quad \mathbf{Ax} = \mathbf{b}$$

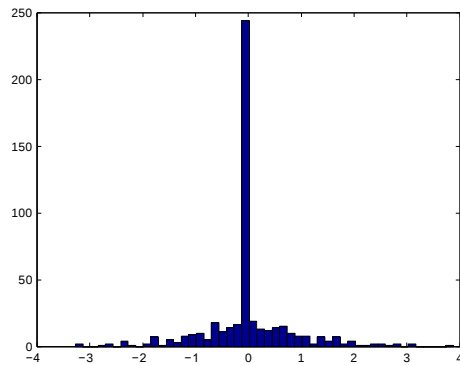
```
X=fmincon(@(x) f(x),zeros(500,1),[],[],A,b,[],[],options);
```

where $f(\mathbf{x})$ is a convex function.

ℓ_1 Norm Regression

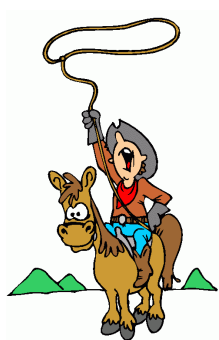


ℓ_2 Solution



ℓ_1 Solution

Least Absolute Shrinkage and Selection Operator (LASSO)



- ▶ LASSO combines shrinking of Ridge regression with variable selection. Tibshirani 1996.
- ▶ Difference between LASSO and Ridge regression is the penalty used

$$\hat{\mathbf{w}}^{\text{ridge}} = \arg \min_{\mathbf{w} \in \mathbb{R}^d} \left[\sum_{i=1}^N (y_i - \sum_{j=0}^d x_{ij} w_j)^2 + \lambda \sum_{j=1}^d w_j^2 \right]$$

$$\hat{\mathbf{w}}^{\text{lasso}} = \arg \min_{\mathbf{w} \in \mathbb{R}^d} \left[\sum_{i=1}^N (y_i - \sum_{j=0}^d x_{ij} w_j)^2 + \lambda \sum_{j=1}^d |w_j| \right]$$

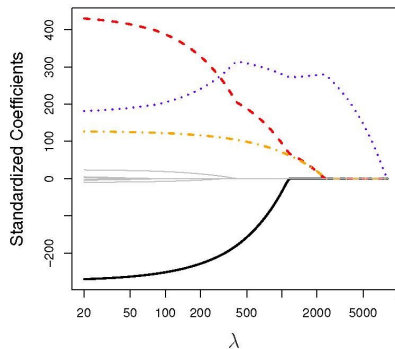
Least Absolute Shrinkage and Selection Operator (LASSO)

- ▶ LASSO coefficients are the solutions to the ℓ_1 optimization problem defined as

$$\begin{aligned}\hat{\mathbf{w}}^{\text{lasso}} &= \arg \min_{\mathbf{w}} \left[\sum_{i=1}^N (y_i - \sum_{j=1}^d x_{ij} w_j)^2 + \lambda \sum_{j=0}^d |w_j| \right] \\ &= \arg \min_{\mathbf{w}} \left[\sum_{i=1}^N (y_i - \mathbf{x}_i^T \mathbf{w})^2 + \lambda \sum_{j=0}^d |w_j| \right] \\ &= \arg \min_{\mathbf{w}} \left[(\mathbf{y} - \mathbf{X}\mathbf{w})^T (\mathbf{y} - \mathbf{X}\mathbf{w}) + \lambda \|\mathbf{w}\|_1 \right].\end{aligned}$$

- ▶ LASSO also shrinks the coefficients.
- ▶ ℓ_1 norm forces coefficients to zero when λ is large: **variable selection**.
- ▶ Lasso yields **sparse** models, keeping subset of variables.
- ▶ Unlike ridge regression, $\hat{\mathbf{w}}_{\lambda}^{\text{lasso}}$ has no closed form.

Lasso Regression Example Credit Data set



- ▶ Lasso performs better when a small number of predictors have strong coefficients, and the remaining predictors are small.
- ▶ Ridge regression performs better when the response is a function of many predictors.

The Variable Selection Property of the Lasso

One can show that the Ridge and Lasso regression coefficient estimates solve the following problems

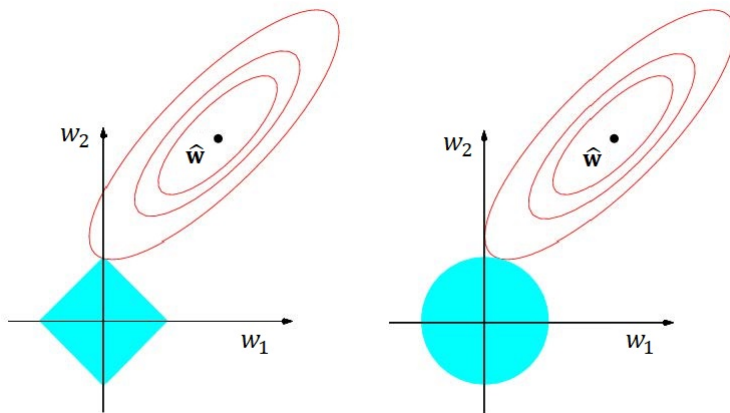
$$\hat{\mathbf{w}}^{\text{ridge}} = \arg \min_{\mathbf{w}} \left\{ \sum_{i=1}^N \left(y_i - \sum_{j=0}^d x_{ij} w_j \right)^2 \right\} \quad (1)$$

$$\text{subject to } \sum_{j=0}^d w_j^2 \leq t$$

$$\hat{\mathbf{w}}^{\text{lasso}} = \arg \min_{\mathbf{w}} \left\{ \sum_{i=1}^N \left(y_i - \sum_{j=0}^d x_{ij} w_j \right)^2 \right\} \quad (2)$$

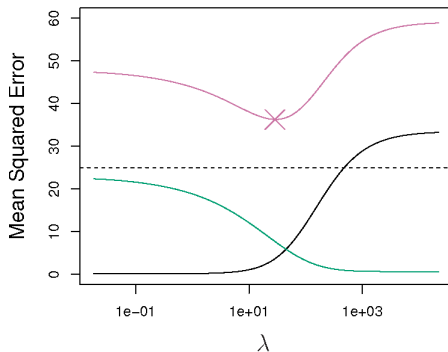
$$\text{subject to } \sum_{j=0}^d |w_j| \leq t$$

The Variable Selection Property of the Lasso



- ▶ RSS has elliptical contours, centered at the LS estimate.
- ▶ Constraint regions, $w_1^2 + w_2^2 \leq t$, and $|w_1| + |w_2| \leq t$.

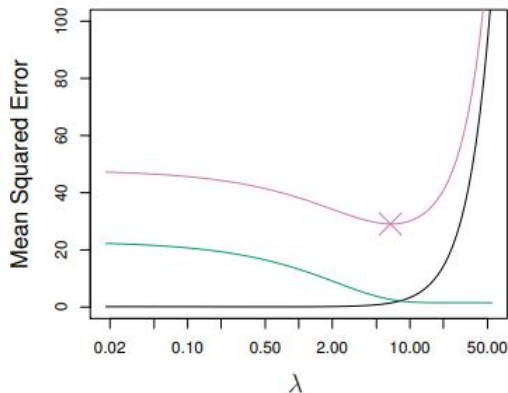
Comparing the Lasso and Ridge Regression



Simulated data set containing $d = 45$ predictors and $n = 50$ observations.
Predictors related to the response.

- Plots of squared bias (black), variance (green), and test MSE (purple) for the lasso.

Comparing the Lasso and Ridge Regression

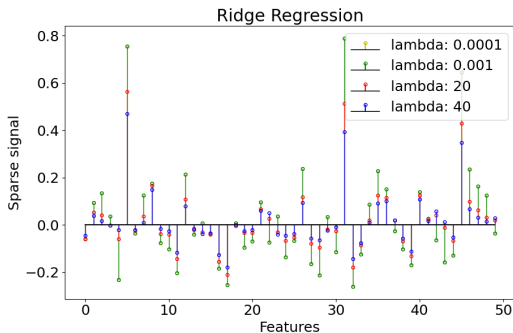
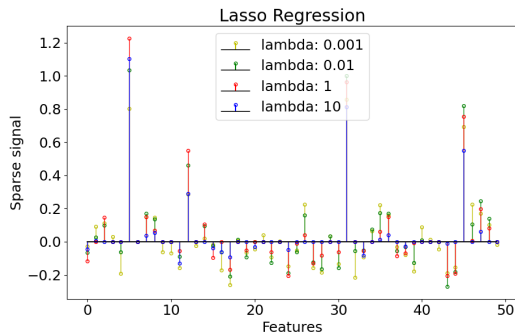


Here the response is a function of only 2 out of 45 predictors.

- Squared bias (black), variance (green), and test MSE (purple) for the lasso.

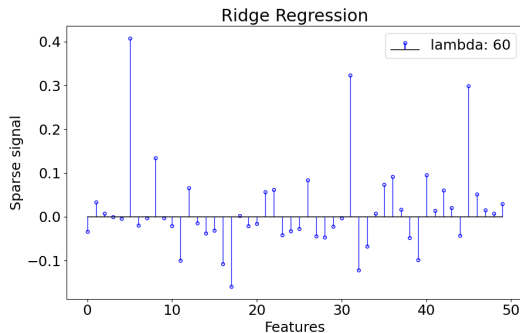
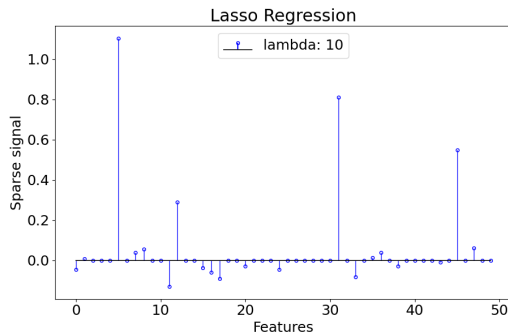
Lasso vs Ridge regression

- ▶ $\mathbf{y} = \mathbf{X}\mathbf{w} + \epsilon$, where $\mathbf{X} \in \mathbb{R}^{40 \times 60}$ is random Gaussian and ϵ is noise.
- ▶ Model given by
$$w(k) = \delta(k - 5) + 0.5\delta(k - 12) + 0.9\delta(k - 31) - 0.75\delta(k - 45)$$

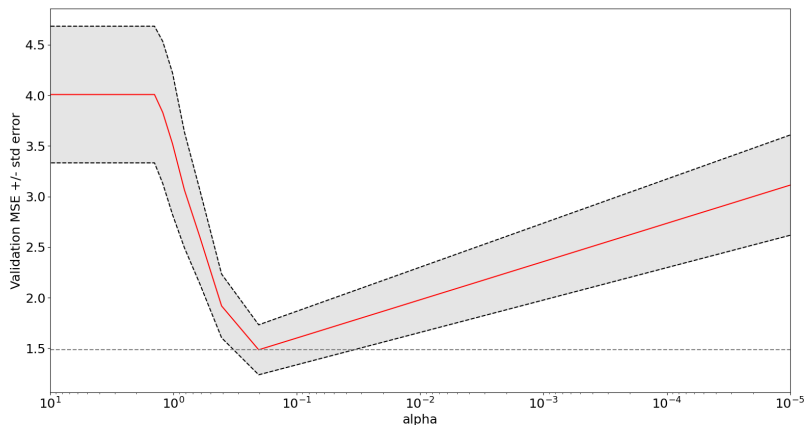


Lasso vs Ridge regression

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- ▶ Model given by
 $w(k) = \delta(k - 5) + 0.5\delta(k - 12) + 0.9\delta(k - 31) - 0.75\delta(k - 45)$



Lasso hyperparameter optimization



Optimization of the alpha parameter through GridSearch with Cross-Validation and Mean Squared Error as the evaluation metric.

Iterative Calculation

- ▶ LASSO does not have a close form solution. Solved iteratively.
- ▶ Define $F(\mathbf{w}) = \|\mathbf{y} - \mathbf{X}\mathbf{w}\|_2^2 + \lambda\|\mathbf{w}\|_1$.
- ▶ The solution to the LASSO problem is denoted as $\mathbf{w}_S$.
- ▶ Define an iterative procedure adding the non-negative term, having zero value at $\mathbf{w}_S$, $G(\mathbf{w}) = (\mathbf{w} - \mathbf{w}_S)^T(\alpha\mathbf{I} - \mathbf{X}^T\mathbf{X})(\mathbf{w} - \mathbf{w}_S)$, to the function $F(\mathbf{w})$.

The cost function is:

$$H(\mathbf{w}) = F(\mathbf{w}) + (\mathbf{w} - \mathbf{w}_S)^T (\alpha \mathbf{I} - \mathbf{X}^T \mathbf{X}) (\mathbf{w} - \mathbf{w}_S), \quad (3)$$

where α is such that the added term is always nonnegative. It means $\alpha > \lambda_{max}$, where λ_{max} is the largest eigenvalue of $\mathbf{X}^T \mathbf{X}$.

$$\begin{aligned} H(\mathbf{w}) &= F(\mathbf{w}) + G(\mathbf{w}) \\ &= \|\mathbf{y} - \mathbf{X}\mathbf{w}\|_2^2 + \lambda \|\mathbf{w}\|_1 + (\mathbf{w} - \mathbf{w}_S)^T (\alpha \mathbf{I} - \mathbf{X}^T \mathbf{X}) (\mathbf{w} - \mathbf{w}_S) \end{aligned}$$

Since $\|\mathbf{w}\|_1 = \mathbf{w}^T \text{sign}\{\mathbf{w}\} = \|w_1 \text{sign}(w_1), w_2 \text{sign}(w_2), \dots, w_N \text{sign}(w_N)\|_1$

$$\begin{aligned} H(\mathbf{w}) &= \|\mathbf{y}\|_2^2 - \mathbf{w}^T \mathbf{X}^T \mathbf{y} - \mathbf{y}^T \mathbf{X} \mathbf{w} + \mathbf{w}^T \mathbf{X}^T \mathbf{X} \mathbf{w} + \lambda \mathbf{w}^T \text{sign}\{\mathbf{w}\} \\ &\quad + (\mathbf{w} - \mathbf{w}_S)^T (\alpha \mathbf{I} - \mathbf{X}^T \mathbf{X}) (\mathbf{w} - \mathbf{w}_S) \end{aligned}$$

Iterative Calculation

$$H(\mathbf{w}) = \|\mathbf{y}\|_2^2 - \mathbf{w}^T \mathbf{X}^T \mathbf{y} - \mathbf{y}^T \mathbf{X} \mathbf{w} + \mathbf{w}^T \mathbf{X}^T \mathbf{X} \mathbf{w} + \lambda \mathbf{w}^T \text{sign}\{\mathbf{w}\} \\ + (\mathbf{w} - \mathbf{w}_S)^T (\alpha \mathbf{I} - \mathbf{X}^T \mathbf{X}) (\mathbf{w} - \mathbf{w}_S)$$

Equating the gradient of $H(\mathbf{w})$ to zero:

$$\begin{aligned} \frac{\partial H(\mathbf{w})}{\partial \mathbf{w}^T} &= -2\mathbf{X}^T \mathbf{y} + 2\mathbf{X}^T \mathbf{X} \mathbf{w} + \lambda \text{sign}\{\mathbf{w}\} + 2(\alpha \mathbf{I} - \mathbf{X}^T \mathbf{X})(\mathbf{w} - \mathbf{w}_S) \\ 0 &= -\mathbf{X}^T \mathbf{y} + \mathbf{X}^T \mathbf{X} \mathbf{w} + \frac{\lambda}{2} \text{sign}\{\mathbf{w}\} + \alpha \mathbf{w} - \mathbf{X}^T \mathbf{X} \mathbf{w} - (\alpha \mathbf{I} - \mathbf{X}^T \mathbf{X}) \mathbf{w}_S \\ 0 &= -\mathbf{X}^T \mathbf{y} + \frac{\lambda}{2} \text{sign}\{\mathbf{w}\} + \alpha \mathbf{w} - (\alpha \mathbf{I} - \mathbf{X}^T \mathbf{X}) \mathbf{w}_S \end{aligned}$$

Rearranging the terms,

$$\mathbf{w} + \frac{\lambda}{2\alpha} \text{sign}\{\mathbf{w}\} = \frac{1}{\alpha} \mathbf{X}^T (\mathbf{y} - \mathbf{X} \mathbf{w}_S) + \mathbf{w}_S$$

Iterative Calculation

Corresponding iterative update

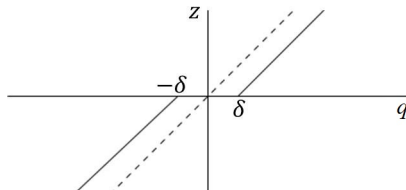
$$\mathbf{w}_{s+1} + \frac{\lambda}{2\alpha} \text{sign}\{\mathbf{w}_{s+1}\} = \frac{1}{\alpha} \mathbf{X}^T (\mathbf{y} - \mathbf{X}\mathbf{w}_s) + \mathbf{w}_s \quad (4)$$

How to solve it?

or

$$z = \text{soft}(q, \delta) = \begin{cases} q + \delta & \text{for } q < -\delta \\ 0 & \text{for } |q| \leq \delta \\ q - \delta & \text{for } q > \delta \end{cases}$$

$$\text{soft}(q, \delta) = \text{sign}(q) \max\{0, |q| - \delta\}$$



Iterative Calculation

► The solution of $z + \delta \text{sign}(z) = q$ is $z = \text{soft}(q, \delta)$

►
$$\underbrace{\mathbf{w}_{s+1}}_z + \underbrace{\frac{\lambda}{2\alpha}}_{\delta} \text{sign} \left\{ \underbrace{\mathbf{w}_{s+1}}_z \right\} = \underbrace{\frac{1}{\alpha} \mathbf{X}^T (\mathbf{y} - \mathbf{X} \mathbf{w}_s) + \mathbf{w}_s}_q$$

Thus,

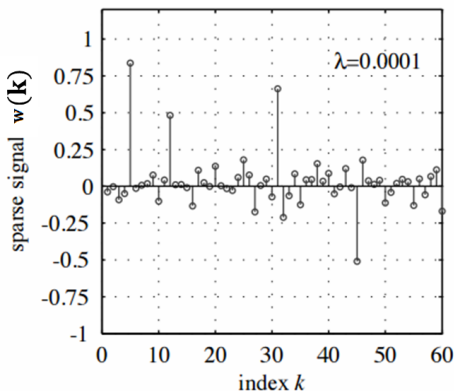
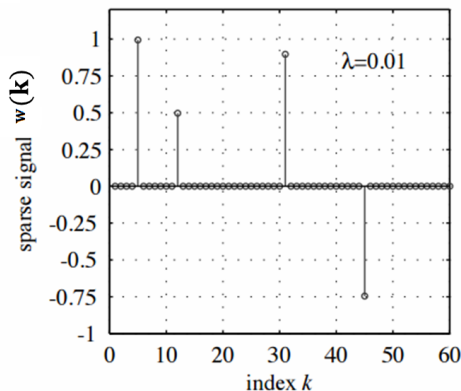
$$\mathbf{w}_{s+1} = \text{soft} \left(\frac{1}{\alpha} \mathbf{X}^T (\mathbf{y} - \mathbf{X} \mathbf{w}_s) + \mathbf{w}_s, \frac{\lambda}{2\alpha} \right) \quad (5)$$

This is the iterative soft-thresholding algorithm (ISTA) for LASSO minimization.

Example

$\mathbf{y} = \mathbf{X}\mathbf{w}$, where

- ▶ $\mathbf{X}$ is a random Gaussian matrix $\in \mathbb{R}^{40 \times 60}$.
- ▶ Oracle model is:
 $w(k) = \delta(k-5) + 0.5\delta(k-12) + 0.9\delta(k-31) - 0.75\delta(k-45)$.
- ▶ The results for $\lambda = 0.01$ and $\lambda = 0.0001$ are presented



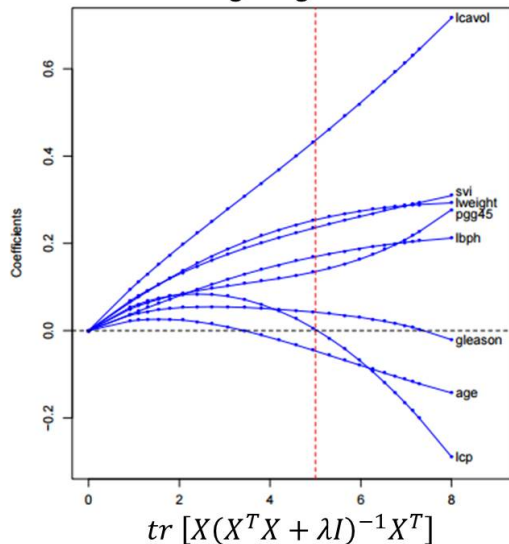
Example: Prostate Cancer

- ▶ Study by Stamey et al. (1989)
- ▶ Examines the correlation between the level of prostate-specific antigen and a number of clinical measures in men who were about to receive radical prostatectomy.

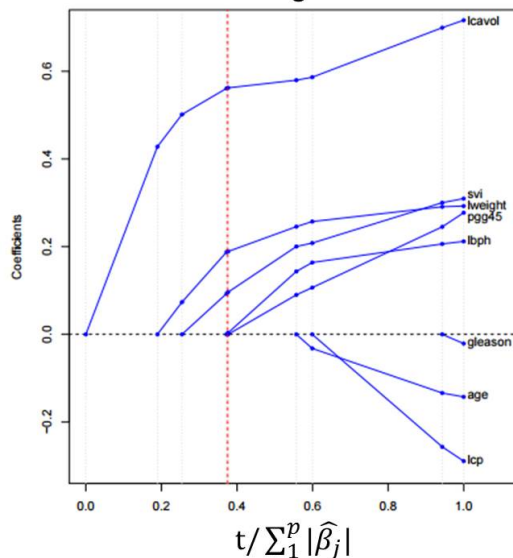
Variable	Unit	Code
Cancer volume	log()	lcavol
Prostate weight	log()	lweight
age	-	age
Amount of benign prostatic hyperlasia	log()	lbph
Seminal Vesicle Invasion	-	svi
Gleason Score	-	gleason
Percentage of Gleason Score	4 or 5	pgg45

Ridge vs Lasso Regression

Ridge Regression



Lasso Regression



Choosing parameters: cross-validation

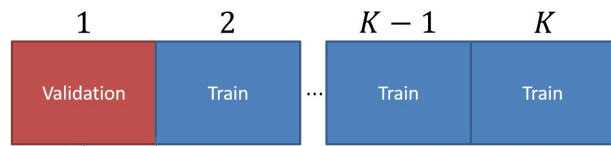
- ▶ Ridge and Lasso have regularization parameters.
- ▶ An *optimal* parameter needs to be chosen in a principled way

K- fold cross-validation: Split data into K equal (or almost equal) parts/folds at random.

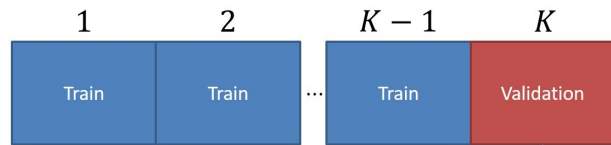
- 1: **for** each value λ_i **do**
- 2: **for** $j = 1, \dots, K$ **do**
- 3: Fit model on data with fold j removed
- 4: Test model on remaining fold j^{th} test error
- 5: **end for**
- 6: Compute average test errors for parameter λ_i
- 7: **end for**
- 8: Pick parameter with smallest average error

Choosing parameters: cross validation

For λ_i



⋮



$w_{\lambda_i}^{j=K}$

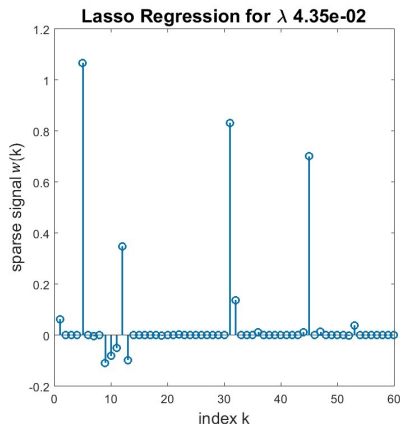
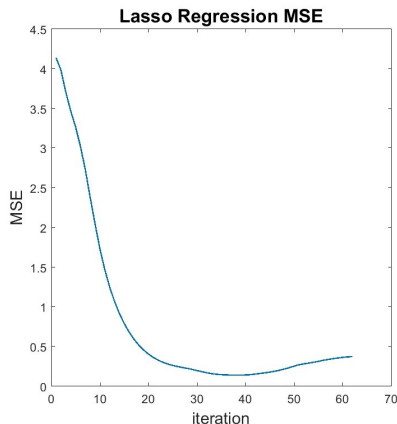
$e_{\lambda_i}^{j_K}$

$$\sum_j e_{\lambda_i}^j = \bar{e}_{\lambda_i}$$

$$\lambda_{opt} = \underset{i,j}{\operatorname{argmin}}(\bar{e}_{\lambda_i})$$

Cross validation- Example K=5

- $\mathbf{y} = \mathbf{X}\mathbf{w} + \epsilon$, where $\mathbf{X} \in \mathbb{R}^{40 \times 60}$ is random Gaussian and ϵ is noise.
- Oracle model is
$$w(k) = \delta(k - 5) + 0.5\delta(k - 12) + 0.9\delta(k - 31) - 0.75\delta(k - 45)$$



Model selection vs Model assesment

- ▶ **Model selection:** estimate performance of different models in order to choose the “best” one
- ▶ **Model assesment:** having a chosen model, estimate its prediction error on new data
- ▶ When enough data is available, it is better to separate the data into three parts: train/validate, and test
- ▶ Typically: 50% train, 25 % validate, 25 % test.
- ▶ Test data is “kept in a vault”, i.e. it is not used to fitting or choosing the model